

Mathematical Institute

Parameter identifiability and model selection for PDE models of cell invasion

YUE LIU *Supervisors: Ruth Baker, Philip Maini Mathematical Institute University of Oxford*

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Oxford Mathematics

Example: what's the best value of *r*?

$$
\frac{\partial C}{\partial t} = rC(1 - C)
$$

- \triangleright Structural identifiability: can the true parameter be recovered, given theoretically *infinite* amount of data? (well-defined, objective property of the model itself)
- ▶ Practical identifiability: can the true parameter be recovered, given a *finite, realistic* amount of data? (subjective, depending on both model and available data)

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$$
AIC = -2\log(p(C_{\text{data}}|\theta)) + 2m \qquad m = # \text{free params}
$$

\n
$$
BIC = -2\log(p(C_{\text{data}}|\theta)) + \log(N)m \qquad N = # \text{data pts}
$$

\n
$$
p(C_{\text{data}}|\theta) = \text{likelihood}
$$

Barrier assay experiments

8 experiments, $n_t = 77$ images, $\Delta t = 20$ min, $n_x = n_y = 150$

How to best model cell invasion?

Spatially-discretized cell density from 2 experiments with different initial conditions [\(link if movie doesn't work\)](http://people.maths.ox.ac.uk/liuy1/figures/barrier_movie.gif)

▶ Are the models practically identifiable?

- ▶ How much data do we need to make the models identifiable?
- ▶ Which model is the "best"?
- ▶ Are parameter estimates consistent across experimental replicates?
- ▶ What's a computationally efficient method to answer these?

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4 candidate models in the competition: $C = C(x, y, t)$

Std. Fisher
$$
(m = 3)
$$
:

\n
$$
\frac{\partial C}{\partial t} = D_0 \nabla^2 C + rC(1 - C/K)
$$
\nPorous Fisher $(m = 4)$:

\n
$$
\frac{\partial C}{\partial t} = \nabla \cdot (D_0(C/K)^{\eta} \nabla C) + rC(1 - C/K)
$$
\nRichards $(m = 4)$:

\n
$$
\frac{\partial C}{\partial t} = D_0 \nabla^2 C + rC(1 - (C/K)^{\gamma})
$$
\nGen. Fisher $(m = 5)$:

\n
$$
\frac{\partial C}{\partial t} = D_0 \nabla^2 C + rC^{\alpha}(1 - C/K)^{\beta}
$$

Line at $log(p) = -1.92$: cutoff for 95% confidence interval Left: identifiable case: smooth, narrow, ∼ parabolic. Right: non-identifiable case: broader, flat top, multimodal, jagged

Yes (but only because we have high resolution data)

To emphasize: these are real, not synthetic, data

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Repeat profile likelihood calculations with temporally down-sampled data (lower *nt*/higher ∆*t*)

The Standard Fisher model remains identifiable even when we down-sample the data to $n_t = 3$

Repeat profile likelihood calculations with temporally down-sampled data (lower *nt*/higher ∆*t*)

The Richards model cease to be identifiable when the data resolution is sufficiently low

Depends on the model.

Std. Fisher (shown), and Porous Fisher: mostly consistent

Depends on the model.

Richards (shown), Generalised Fisher: much less consistent

- ▶ The 4 models (Std Fisher, Porous Fisher, Richards, Gen Fisher) are all identifiable given high resolution data
- ▶ Richards and Gen Fisher becomes non-identifiable if data resolution is low
- \triangleright These two models also show inconsistencies across experiment replicates even if they appears identifiable
- ▶ **Inconsistency reflects sensitivity of the model to process noise, a symptom of non-identifiability**
- ▶ Experimental initial conditions can have a major impact on parameter identifiability
- ▶ Computational cost of profile likelihood compares favourably to MCMC, while still being informative

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θ: model parameters, θ−*ⁱ* : parameters except θ*ⁱ*

Maximum likelihood estimate (MLE):

$$
\boldsymbol{\theta}^* = \arg\!\sup_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|C_{\text{data}})
$$

Bayesian inference:

$$
p(\boldsymbol{\theta}|C_{\text{data}}) \sim p(C_{\text{data}}|\boldsymbol{\theta})p(\boldsymbol{\theta})
$$

Profile likelihood:

$$
p(\theta_i = \theta'_i | C_{data}) \sim \max_{\theta_{-i}} p(C_{data} | \theta_{-i}, \theta_i = \theta'_i)
$$

Triangular initial condition makes the parameter less consistent. Parameter estimates in Richards and Gen Fisher models can be far from the estimates with circular initial conditions

Result: Does profile likelihood agree with MCMC?

Yes

The samples generated from Metropolis-Hastings MCMC closely matches the contours of the two-parameter profile likelihood function

We measure computational cost by the number of model simulations required to compute the profile likelihood curves (all other costs negligible)

Total cost \approx # free parameters $*$ # points per curve (we chose 10) $*$ average $#$ model simulations needed for optimization

- ▶ 3 f.p. (Standard Fisher): 3 * 10 * (40 60) ≈ 1200 − 1800
- ▶ 4 f.p. (Porous Fisher, Richards): $4 * 10 * (60 100)$ \approx 2400 $-$ 4000

▶ 5 f.p. (Gen. Fisher): 5 * 10 * (140 – 250 ≈ 7000 − 12500 Optimization may fail with off-the-shelf methods with $>$ 5 f.p.